

**German-Japanese conference on  
Stochastic Analysis and Applications**

**Continuity and bounds of the transition  
density functions of the solutions to  
path-dependent stochastic differential  
equations**

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## **0. Acknowledgment**

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# 1. Introduction

Recently I study, by means of Stochastic Analysis, the Gaussian two-sided bounds of the fund. solution and the Hölder (Lipschitz) continuity of the solution to the parabolic partial differential equation:

$$\left\{ \begin{array}{l} \frac{\partial}{\partial t} u(t, x) = \frac{1}{2} \sum_{i,j=1}^d a_{ij}(t, x) \frac{\partial^2}{\partial x_i \partial x_j} u(t, x) \\ \quad + \sum_{i=1}^d b_i(t, x) \frac{\partial}{\partial x_i} u(t, x) + c(t, x) u(t, x) \\ u(0, x) = f(x) \end{array} \right.$$

where  $a, b, c$  have bad regularities

and  $a$  is uniformly positive definite.

In this talk we consider stochastic differential equations with path-dependent drift terms.

Consider the path-dependent SDE:

$$\begin{cases} dY_t^x &= \sigma(t, Y_t^x)dB_t + b(t, Y_t^x)dt + \tilde{b}(t, Y^x)dt \\ Y_0^x &= x. \end{cases}$$

where

$\sigma: [0, T] \times \mathbb{R}^d \rightarrow \mathbb{R}^d \times \mathbb{R}^d$  , measurable

$b: [0, T] \times \mathbb{R}^d \rightarrow \mathbb{R}^d$  , measurable

$\tilde{b}: [0, T] \times C([0, T]; \mathbb{R}^d) \rightarrow \mathbb{R}^d$  such that

$\mathcal{B}([0, T]) \otimes \mathcal{B}(C([0, T]; \mathbb{R}^d)) / \mathcal{B}(\mathbb{R}^d)$ -measurable

Let

$$\pi_t: C([0, T]; \mathbb{R}^d) \rightarrow C([0, T]; \mathbb{R}^d)$$

defined by  $(\pi_t w) = w(\cdot \wedge t)$

$$\mathcal{B}_t(C([0, T]; \mathbb{R}^d)) := \pi_t^{-1}[\mathcal{B}(C([0, T]; \mathbb{R}^d))]$$

## Assumption

$\tilde{b}(t, \cdot): \mathcal{B}_t(C([0, T]; \mathbb{R}^d)) / \mathcal{B}(\mathbb{R}^d)$ -m'ble  $\forall t \in [0, T]$ ,

$$\|\tilde{b}\|_\infty := \sup_{(t, w) \in [0, T] \times C([0, T]; \mathbb{R}^d)} |\tilde{b}(t, w)| < \infty.$$

Let  $a(t, x) := \sigma(t, x)^T \sigma(t, x)$  for  $(t, x) \in [0, T] \times \mathbb{R}^d$ .

## Assumption

$$\Lambda^{-1}I \leq a(t, x) \leq \Lambda I, \quad (t, x) \in [0, T] \times \mathbb{R}^d,$$

$\exists M_a$ : a conti. func. on  $[0, \infty)$  such that  $M_a(0) = 0$ ,

$$\sup_{t \in [0, T]} |a(t, x) - a(t, y)| \leq M_a(|x - y|),$$

$$\int_0^1 \frac{1}{\tilde{r}} \left( \int_0^{\tilde{r}} \frac{1}{r} M_a(r) dr \right) d\tilde{r} < \infty,$$

$$\|b\|_\infty := \sup_{(t, x) \in [0, T] \times \mathbb{R}^d} |b(t, x)| < \infty$$

$\exists M_b$ : a conti. func. on  $[0, \infty)$  such that  $M_b(0) = 0$ ,

$$\sup_{t \in [0, T]} |b(t, x) - b(t, y)| \leq M_b(|x - y|), \quad \int_0^1 \frac{1}{r} M_b(r) dr < \infty.$$

Consider the SDE of Markov type:

$$\begin{cases} dX_t^x = \sigma(t, X_t^x)dB_t + b(t, X_t^x)dt \\ X_0^x = x. \end{cases}$$

Then, the transition probability density function  $q(s, x; t, y)$  of  $X$  and the estimates

$$\begin{aligned} \frac{C_-}{(t-s)^{d/2}} \exp\left(-\frac{\gamma_- |x-y|^2}{t-s}\right) \\ \leq q(s, x; t, y) \leq \frac{C_+}{(t-s)^{d/2}} \exp\left(-\frac{\gamma_+ |x-y|^2}{t-s}\right), \\ |\nabla_x q(s, x; t, y)| \leq \frac{C_+}{(t-s)^{(d+1)/2}} \exp\left(-\frac{\gamma_+ |x-y|^2}{t-s}\right) \end{aligned}$$

for  $(s, x), (t, y) \in [0, T] \times \mathbb{R}^d$  s.t.  $s < t$ .

These are obtained by parametrix.

## 2. Perturbation

$$\mathcal{E}(t, X^x) := \exp \left( \int_0^t \langle \tilde{b}_\sigma(s, X^x), dB_s \rangle - \frac{1}{2} \int_0^t |\tilde{b}_\sigma(s, X^x)|^2 ds \right)$$

where  $\tilde{b}_\sigma(t, w) := \sigma(t, w_t)^{-1} \tilde{b}(t, w)$ .

Then, by the Girsanov transform

$$E [f(Y_t^x)] = E [f(X_t^x) \mathcal{E}(t, X^x)].$$

We regard  $Y$  as a perturbation of  $X$ ,

and consider estimates of the density of  $Y_t$ .



We have the existence of the density function  $p(0, x; t, y)$  of  $Y_t^x$  and

$$p(0, x; t, y) = q(0, x; t, y) E^{X_t^x=y} [\mathcal{E}(t, X^x)]$$

(for a.e.  $y$  w.r.t. the Lebesgue measure)

where  $E^{X_t^x=y}[\cdot]$  is the expectation w.r.t. the regular conditional probability measure given by  $X_t^x = y$ .

We have known that  $q(s, x; t, y)$  satisfies the Gaussian estimates.

Hence, to see the Gaussian estimates of  $p(0, x; t, y)$  it is sufficient to have estimates of  $E^{X_t^x=y} [\mathcal{E}(t, X^x)]$ .

Let  $p^s(0, x; t, y) := E [\mathcal{E}(s, X^x)q(s, X_s^x; t, y)]$   
for  $s \in [0, t)$  and  $x, y \in \mathbb{R}^d$ .

By the Markov property we have

$$E [\mathcal{E}(s, X^x)q(s, X_s^x; t, y)] = q(0, x; t, y)E^{X_t^x=y} [\mathcal{E}(s, X^x)]$$

for  $s \in [0, t)$ .

**Lemma** For each  $r \in \mathbb{R}$ ,

$$\sup_{s \in [0, t)} E [\mathcal{E}(s, X^x)^r q(s, X_s^x; t, y)] \leq Ct^{-d/2} \exp \left( -\frac{\gamma|x-y|^2}{t} \right)$$

**Proof.** Since

$$\begin{aligned}
& -\frac{\partial}{\partial s}q(s, x; t, y) = \frac{\partial}{\partial t}q(s, x; t, y) \\
& = \frac{1}{2} \sum_{i,j=1}^d a_{ij}(t, x) \frac{\partial^2}{\partial x_i \partial x_j} q(s, x; t, y) + \sum_{i=1}^d b_i(t, x) \frac{\partial}{\partial x_i} q(s, x; t, y),
\end{aligned}$$

by Itô's formula we have

$$\begin{aligned}
& E[\mathcal{E}(s, X^x)^r q(s, X_s^x; t, y)] - q(0, x; t, y) \\
& = \frac{r(r-1)}{2} \int_0^s E \left[ \mathcal{E}(u, X^x)^r q(u, X_u^x; t, y) |\tilde{b}_\sigma(u, X)|^2 \right] du \\
& \quad + \frac{r}{2} \int_0^s E \left[ \mathcal{E}(u, X^x)^r \langle [\nabla q(u, \cdot; t, y)](X_u^x), \tilde{b}(u, X) \rangle \right] du.
\end{aligned}$$

□

Hence, we have the following.

**Lemma**  $p(0, x; t, y) = \lim_{s \nearrow t} p^s(0, x; t, y)$

**Proposition** For each  $r \in \mathbb{R}$ ,

$$q(0, x; t, y) E^{X_t^x=y} [\mathcal{E}(t, X^x)^r] \leq C t^{-d/2} \exp\left(-\frac{\gamma|x-y|^2}{t}\right)$$

Since  $p(0, x; t, y) = q(0, x; t, y) E^{X_t^x=y} [\mathcal{E}(t, X^x)]$ ,

we obtain the Gaussian upper bound of  $p(0, x; t, y)$ .

Since

$$\begin{aligned} 1 &= E^{X_t^x=y} \left[ \mathcal{E}(t, X^x)^{1/2} \mathcal{E}(t, X^x)^{-1/2} \right]^2 \\ &\leq E^{X_t^x=y} [\mathcal{E}(t, X^x)] E^{X_T^x=y} [\mathcal{E}(t, X^x)^{-1}], \end{aligned}$$

we have

$$E^{X_t^x=y} [\mathcal{E}(t, X^x)^{-1}]^{-1} \leq E^{X_t^x=y} [\mathcal{E}(t, X^x)].$$

Hence,

$$\begin{aligned} p(0, x; t, y) &= q(0, x; t, y) E^{X_t^x=y} [\mathcal{E}(t, X^x)] \\ &\geq \frac{q(0, x; t, y)^2}{q(0, x; t, y) E^{X_t^x=y} [\mathcal{E}(t, X^x)^{-1}]} \\ &\geq \frac{C_-^2 t^{-d} \exp(-2\gamma_- |x - y|^2/t)}{C t^{-d/2} \exp(-\gamma |x - y|^2/t)} \end{aligned}$$

## Theorem

There exist positive constants  $\tilde{C}_-$ ,  $\tilde{\gamma}_-$ ,  $\tilde{C}_+$  and  $\tilde{\gamma}_+$  depending on  $T$ ,  $d$ ,  $\Lambda$ ,  $\|b\|_\infty$ ,  $\|\tilde{b}\|_\infty$ ,  $M_a$  and  $M_b$  such that

$$\begin{aligned} \tilde{C}_- t^{-d/2} \exp\left(-\frac{\tilde{\gamma}_- |x-y|^2}{t}\right) \\ \leq p(0, x; t, y) \leq \tilde{C}_+ t^{-d/2} \exp\left(-\frac{\tilde{\gamma}_+ |x-y|^2}{t}\right) \end{aligned}$$

for  $t \in [0, T]$  and  $x, y \in \mathbb{R}^d$ .

Next we consider the continuity of

the density function.

**Theorem** Let  $t \in (0, T]$  be fixed. Assume that

$\exists \beta \in [0, 1]$  and a conti. func.  $\rho$  s.t.  $\rho(0) = 0$ ,

$$|q(s, x; t, y_1) - q(s, x; t, y_2)| \leq (t - s)^{-(d+\beta)/2} \rho(|y_1 - y_2|)$$

$$\begin{aligned} |\nabla_x q(s, x; t, y_1) - \nabla_x q(s, x; t, y_2)| \\ \leq (t - s)^{-(d+2\beta)/2} \rho(|y_1 - y_2|) \end{aligned}$$

for  $s \in [0, t)$  and  $x, y_1, y_2 \in \mathbb{R}^d$ . Then, for any  $\varepsilon \in (0, 1)$ ,

$$\begin{aligned} & |p(0, x; t, y_1) - p(0, x; t, y_2)| \\ & \leq \tilde{C} t^{-[d+\beta(1-\varepsilon)]/2} \rho(|y_1 - y_2|)^{1-\varepsilon} \\ & \quad \times \exp\left(-\frac{\tilde{\gamma} \min\{|x - y_1|^2, |x - y_2|^2\}}{t}\right). \end{aligned}$$

**Remark:** Consider the case that  $X^x$  is a  $d$ -dimensional standard Brownian motion. In this case,

$$q(s, x; t, y) = (2\pi)^{-d/2} (t - s)^{-d/2} \exp\left(-\frac{|x - y|^2}{2(t - s)}\right).$$

Then,

$$|q(s, x; t, y_1) - q(s, x; t, y_2)| \leq C(t - s)^{-(d+1)/2} |y_1 - y_2|$$

$$|\nabla_x q(s, x; t, y_1) - \nabla_x q(s, x; t, y_2)| \leq C'(t - s)^{-(d+2)/2} |y_1 - y_2|.$$

Hence, we let  $\beta := 1$  and  $\rho(r) := Cr$ .



**Proof.** By the interpolation inequality,

$$\begin{aligned}
& \left| \nabla_{\xi} q(s, \xi; t, y_1) - \nabla_{\xi} q(s, \xi; t, y_2) \right| \\
& \leq C(t-s)^{-d/2-\beta(1-\varepsilon)} \rho(|y_1 - y_2|)^{1-\varepsilon} \\
& \quad \times \max_{\eta=y_1, y_2} \exp\left(-\frac{\varepsilon\gamma_+ |\xi - \eta|^2}{t-s}\right) \\
& \left| q(0, x; t, y_1) - q(0, x; t, y_2) \right| \\
& \leq Ct^{-d/2-\beta(1-\varepsilon)/2} \rho(|y_1 - y_2|)^{1-\varepsilon} \max_{\eta=y_1, y_2} \exp\left(-\frac{\varepsilon\gamma_+ |\xi - \eta|^2}{t}\right)
\end{aligned}$$

$$\begin{aligned}
& |p^s(0, x; t, y_1) - p^s(0, x; t, y_2)| \\
&= |E[\mathcal{E}(s, X^x)q(s, X_s^x; t, y_1)] - E[\mathcal{E}(s, X^x)q(s, X_t^x; t, y_2)]| \\
&\leq |q(0, x; t, y_1) - q(0, x; t, y_2)| \\
&\quad + \frac{1}{2} \left| \int_0^s E[\mathcal{E}(u, X^x) \langle [\nabla q(u, \cdot; t, y_1)(X_u^x) \right. \\
&\quad \quad \quad \left. - \nabla q(u, \cdot; t, y_2)(X_u^x)], b(u, X) \rangle] du \right| \\
&\leq Ct^{-d/2-\beta(1-\varepsilon)/2} \rho(|y_1 - y_2|)^{1-\varepsilon} \max_{\eta=y_1, y_2} \exp\left(-\frac{\varepsilon\gamma_+ |x - \eta|^2}{t}\right) \\
&\quad + C\rho(|y_1 - y_2|)^{1-\varepsilon} \int_0^s (t-u)^{-d/2-\beta(1-\varepsilon)} \\
&\quad \quad \times E\left[\mathcal{E}(u, X^x) \max_{\eta=y_1, y_2} \exp\left(-\frac{\varepsilon\gamma_+ |X_u^x - \eta|^2}{t-u}\right)\right] du
\end{aligned}$$

□

### 3. Path-dependent drift perturbation of symmetric diffusion processes

$\sigma \in C_b^1(\mathbb{R}^d; \mathbb{R}^d \otimes \mathbb{R}^d)$  s.t.

$\exists M_\sigma \in C([0, \infty); \mathbb{R}_+)$  satisfying  $M_\sigma(0) = 0$ ,

$$|\nabla\sigma(x) - \nabla\sigma(y)| \leq M'_\sigma(|x - y|), \text{ and } \int_0^1 \frac{M'_\sigma(r)}{r} dr < \infty.$$

Assume  $\Lambda^{-1}I \leq \sigma(x)\sigma(x)^T \leq \Lambda I$ .

Define  $b \in C_b(\mathbb{R}^d; \mathbb{R}^d)$  by

$$\begin{aligned} b_i(x) &:= \frac{1}{2} \sum_{j=1}^d \frac{\partial}{\partial x_j} \left( \sum_{k=1}^d \sigma_{jk}(x)\sigma_{ik}(x) \right) \\ &= \frac{1}{2} \sum_{j,k=1}^d \left( \frac{\partial\sigma_{jk}}{\partial x_j}(x)\sigma_{ik}(x) + \sigma_{jk}(x)\frac{\partial\sigma_{ik}}{\partial x_j}(x) \right). \end{aligned}$$

Consider

$$\begin{cases} dX_t^x = \sigma(X_t^x)dB_t + b(X_t^x)dt \\ X_0^x = x. \end{cases}$$

The results by parametrix imply that the transition probability density function  $q(t, x, y)$  exists and

$$\frac{C_-}{t^{d/2}} \exp\left(-\frac{\gamma_- |x - y|^2}{t}\right) \leq q(t, x, y) \leq \frac{C_+}{t^{d/2}} \exp\left(-\frac{\gamma_+ |x - y|^2}{t}\right)$$

$$|\nabla_x q(t, x, y)| \leq Ct^{-(d+1)/2} \exp\left(-\frac{\gamma |x - y|^2}{t}\right)$$

$$|\nabla_x^2 q(t, x, y)| \leq Ct^{-(d+2)/2} \exp\left(-\frac{\gamma |x - y|^2}{t}\right)$$

for  $t \in [0, T]$  and  $x, y \in \mathbb{R}^d$ .

Since

$$\begin{aligned} & \frac{1}{2} \sum_{i,j,k=1}^d \sigma_{ik}(x) \sigma_{jk}(x) \frac{\partial^2}{\partial x_i \partial x_j} + \frac{1}{2} \sum_{i=1}^d b_i(x) \frac{\partial}{\partial x_i} \\ &= \frac{1}{2} \sum_{i,j=1}^d \frac{\partial}{\partial x_j} \left( \sum_{k=1}^d \sigma_{jk}(x) \sigma_{ik}(x) \right) \frac{\partial}{\partial x_i}, \end{aligned}$$

the generator is symmetric on  $L^2(\mathbb{R}^d, dx)$ . Hence,

$$q(t, x, y) = q(t, y, x).$$

Therefore,

$$|\nabla_x \nabla_y q(t, x, y)| \leq C t^{-(d+2)/2} \exp\left(-\frac{\gamma|x-y|^2}{t}\right).$$

Let  $\tilde{b}$  be an  $[0, T] \times C([0, T]; \mathbb{R}^d) \rightarrow \mathbb{R}^d$  s.t.

$\mathcal{B}([0, T]) \otimes \mathcal{B}(C([0, T]; \mathbb{R}^d)) / \mathcal{B}(\mathbb{R}^d)$ -measurable

and  $\tilde{b}(t, \cdot)$  is  $\mathcal{B}_t(C([0, T]; \mathbb{R}^d)) / \mathcal{B}(\mathbb{R}^d)$ -measurable.

Consider the path-dependent stochastic differential equation

$$\begin{cases} dY_t^x &= \sigma(t, Y_t^x)dB_t + \tilde{b}(t, Y^x)dt \\ Y_0^x &= x. \end{cases}$$

We assume

$$\|\tilde{b}\|_\infty := \sup_{(t, w) \in [0, T] \times C([0, T]; \mathbb{R}^d)} |\tilde{b}(t, w)| < \infty.$$

## Corollary

Then, the distribution of  $Y_t^x$  has the density function  $p(0, x; t, \cdot)$  which satisfies

$$\begin{aligned} & \tilde{C}_- t^{-d/2} \exp\left(-\frac{\tilde{\gamma}_- |x - y|^2}{t}\right) \\ & \leq p(0, x; t, y) \leq \tilde{C}_+ t^{-d/2} \exp\left(-\frac{\tilde{\gamma}_+ |x - y|^2}{t}\right) \end{aligned}$$

and for any  $\varepsilon \in (0, 1]$

$$\begin{aligned} & |p(0, x; t, y_1) - p(0, x; t, y_2)| \\ & \leq \tilde{C} t^{-(d+1)/2} |y_1 - y_2|^{1-\varepsilon} \exp\left(-\frac{\tilde{\gamma} \min\{|x - y_1|^2, |x - y_2|^2\}}{t}\right). \end{aligned}$$

## 4. Further consideration to less regular diffusion coefficients

$$\begin{cases} dX_t^x = \sigma(t, X_t^x)dB_t, & t \in [0, T] \\ X_0^x = x. \end{cases}$$

where  $\sigma : [0, T] \times \mathbb{R}^d \rightarrow \mathbb{R}^d \otimes \mathbb{R}^d$ .

Assume  $\sigma$  is smooth and consider a priori estimates.

Denote the transition probability density function of  $X$  by  $q(s, x; t, y)$ .

Let  $a(t, x) := \sigma(t, x)\sigma(t, x)^T$  for  $(t, x) \in [0, T] \times \mathbb{R}^d$ .



We assume

$$\begin{aligned} \Lambda^{-1}I &\leq a(t, x) \leq \Lambda I, \quad (t, x) \in [0, \infty) \times \mathbb{R}^d, \\ \frac{C_-}{(t-s)^{d/2}} \exp\left(-\frac{\gamma_- |x-y|^2}{t-s}\right) \\ &\leq q(s, x; t, y) \leq \frac{C_+}{(t-s)^{d/2}} \exp\left(-\frac{\gamma_+ |x-y|^2}{t-s}\right). \end{aligned}$$

Moreover, we assume

$$\sum_{i,j=1}^d \sup_{t \in [0, T]} \int_{\mathbb{R}^d} \left| \frac{\partial}{\partial x_j} a_{ij}(s, x) \right|^\theta e^{-m|x|} dx \leq M$$

where  $\theta$  is a constant in  $[d, \infty) \cap (2, \infty)$ ,  $m$  and  $M$  are nonnegative constants.

We consider the perturbation by path-dependent drift.

Let  $\tilde{b} : [0, T] \times C([0, T]; \mathbb{R}^d) \rightarrow \mathbb{R}^d$  s.t.

$\tilde{b}$  is  $\mathcal{B}([0, T]) \otimes \mathcal{B}(C([0, T]; \mathbb{R}^d)) / \mathcal{B}(\mathbb{R}^d)$ -measurable  
and  $\tilde{b}(t, \cdot)$  is  $\mathcal{B}_t(C([0, T]; \mathbb{R}^d)) / \mathcal{B}(\mathbb{R}^d)$ -measurable.

Consider the path-dependent stochastic differential equation

$$\begin{cases} dY_t^x = \sigma(t, Y_t^x)dB_t + \tilde{b}(t, Y^x)dt \\ Y_0^x = x. \end{cases}$$

We assume the boundedness of  $\tilde{b}$ . As we have seen, there exists a density function  $p(0, x; t, \cdot)$  of  $Y_t^x$ .

Define  $\mathcal{E}(t, X^x)$  and  $p^s(0, x; t, y)$  as before.

Now we see the Gaussian two-sided bounds of  $Y_t^x$ .  
 The disadvantage of the setting in this section is that  
 we do not have the estimate

$$|\nabla_x q(s, x; t, y)| \leq \frac{C_+}{(t-s)^{(d+1)/2}} \exp\left(-\frac{\gamma_+ |x-y|^2}{t-s}\right).$$

A problem appears, when we estimate

$$\left| \int_0^t E \left[ \mathcal{E}(s, X^x)^r \langle [\nabla q(s, \cdot; t, y)](X_s^x), \tilde{b}(s, X) \rangle \right] ds \right|.$$

$$\begin{aligned}
& \int_{t_1}^{t_2} E \left[ \mathcal{E}(s, X^x)^r \left| \langle [\nabla q(s, \cdot; t, y)](X_s^x), \tilde{b}(s, X) \rangle \right| \right] ds \\
& \leq \|\tilde{b}\|_\infty \int_{t_1}^{t_2} E \left[ \mathcal{E}(s, X^x)^r q(s, X_s^x; t, y)^{1/4} \right. \\
& \quad \times \left( [s(t-s)]^{-1/2} q(s, X_s^x; t, y) \right)^{1/4} \\
& \quad \left. \times [s(t-s)]^{1/8} \frac{|[\nabla q(s, \cdot; t, y)](X_s^x)|}{q(s, X_s^x; t, y)^{1/2}} \right] ds \\
& \leq \|\tilde{b}\|_\infty \left( \int_{t_1}^{t_2} E \left[ \mathcal{E}(s, X^x)^{4r} q(s, X_s^x; t, y) \right] ds \right)^{1/4} \\
& \quad \times \left( \int_{t_1}^{t_2} [s(t-s)]^{-1/2} E \left[ q(s, X_s^x; t, y) \right] ds \right)^{1/4} \\
& \quad \times \left( \int_{t_1}^{t_2} [s(t-s)]^{1/4} E \left[ \frac{|[\nabla q(s, \cdot; t, y)](X_s^x)|^2}{q(s, X_s^x; t, y)} \right] ds \right)^{1/2}
\end{aligned}$$

To estimate

$$\int_0^t [s(t-s)]^{1/4} \int_{\mathbb{R}^d} \frac{|\nabla_{\xi} q(s, \xi; t, y)|^2}{q(s, \xi; t, y)} q(0, x; s, \xi) d\xi ds$$

we use the estimate

$$\begin{aligned} & \int_0^t [s(t-s)]^{1/4} \int_{\mathbb{R}^d} \frac{|\nabla_{\xi} q(s, \xi; t, y)|^2}{q(s, \xi; t, y)} q(0, x; s, \xi) d\xi ds \\ & \leq Ct^{-d/2+1/2} (1 + |\log t|). \end{aligned}$$

By this estimate we have

$$\begin{aligned} & \int_0^t E \left[ \mathcal{E}(s, X^x)^r \left| \langle [\nabla q(s, \cdot; t, y)](X_s^x), \tilde{b}(s, X) \rangle \right| \right] ds \\ & \leq Ct^{-d/2} \exp \left( -\frac{\gamma|x-y|^2}{t} \right). \end{aligned}$$

## Proposition

There exist positive constants  $\tilde{C}_-$ ,  $\tilde{\gamma}_-$ ,  $\tilde{C}_+$  and  $\tilde{\gamma}_+$  depending on  $T$ ,  $r$ ,  $d$ ,  $\Lambda$ ,  $\|\tilde{b}\|_\infty$ ,  $C_-$ ,  $\gamma_-$ ,  $C_+$ ,  $\gamma_+$ ,  $\theta$ ,  $m$  and  $M$ , such that

$$\begin{aligned} & \tilde{C}_- t^{-d/2} \exp\left(-\frac{\tilde{\gamma}_- |x - y|^2}{t}\right) \\ & \leq p(0, x; t, y) \leq \tilde{C}_+ t^{-d/2} \exp\left(-\frac{\tilde{\gamma}_+ |x - y|^2}{t}\right) \end{aligned}$$

for  $t \in (0, T]$  and  $x, y \in \mathbb{R}^d$ .

To obtain the Gaussian two-sided bound of the limit process, we perhaps need the continuity of  $\tilde{b}$ .

Thank you for your attention!

## 5. Remark in the case of starting at $t_0$

Let  $t_0 \in [0, T)$  and  $w \in C([0, t_0]; \mathbb{R}^d)$ . Consider

$$\begin{cases} dZ_t^{(t_0, w)} = \sigma(t, Z_t^{(t_0, w)})dB_t + b(t, Z_t^{(t_0, w)})dt + \tilde{b}(t, Z_t^{(t_0, w)})dt \\ Z_s^{(t_0, w)} = w_s, \quad s \in [0, t_0]. \end{cases}$$

We consider a reduction to the previous case. For

$w \in C([0, t_0]; \mathbb{R}^d)$  and  $\tilde{w} \in C([0, T - t_0]; \mathbb{R}^d)$

define  $w \bowtie_{t_0} \tilde{w} \in C([0, T]; \mathbb{R}^d)$  by

$$(w \bowtie_{t_0} \tilde{w})_t := \begin{cases} w_t, & t \in [0, t_0] \\ \tilde{w}_{t_0+t}, & t \in (t_0, T]. \end{cases}$$



Let  $\tilde{b}^{(t_0, w)}(t, \tilde{w}) := \tilde{b}(t + t_0, w \bowtie_{t_0} \tilde{w})$

for  $(t, \tilde{w}) \in [t_0, T] \times C([0, T - t_0]; \mathbb{R}^d)$  such that  $\tilde{w}_0 = w_{t_0}$ .

Let  $\sigma^{t_0}(t, x) := \sigma(t_0 + t, x)$  and  $b^{t_0}(t, x) := b(t_0 + t, x)$ .

Let  $Y^{w_{t_0}}$  be the solution to

$$\begin{cases} dY_t^{w_{t_0}} = \sigma^{t_0}(t, Y_t^{w_{t_0}})dB_t^{t_0} + b^{t_0}(t, Y_t^{w_{t_0}})dt \\ \quad \quad \quad \quad \quad + \tilde{b}^{(t_0, w)}(t, Y^{w_{t_0}})dt, & t \in [0, T - t_0] \\ Y_0^{w_{t_0}} = w_{t_0}. \end{cases}$$

where  $B^{t_0} = B_{t_0 + \cdot}$ .

Then, we have

$$Z^{(t_0, w)}(t_0 + t) = Y^{w_{t_0}}(t) \text{ for } t \in [0, T - t_0].$$

Therefore, we obtain the same conclusion in this case.

## 6. Remark on unbounded drift terms.

Consider the one-dimensional SDE

$$\begin{cases} dX_t^x = dB_t + \kappa X_t^x dt, & t \in [0, \infty) \\ X_0^x = x. \end{cases}$$

where  $\kappa \in \mathbb{R} \setminus \{0\}$ . The solution  $X^x$  is explicitly written as

$$X_t^x = e^{\kappa t} \left( x + \int_0^t e^{-\kappa s} dB_s \right), \quad t \in [0, \infty).$$

Hence, the distribution of  $X_t^x$  is

$$\sqrt{\frac{e^{2\kappa t} - 1}{4\pi\kappa}} \exp\left(-\frac{\kappa(y - xe^{\kappa t})^2}{e^{2\kappa t} - 1}\right) dy.$$

Since

$$\exp\left(-\frac{\kappa(y - xe^{\kappa t})^2}{e^{2\kappa t} - 1}\right)\Big|_{y=x} = \exp\left(-\frac{\kappa x^2(1 - e^{\kappa t})^2}{e^{2\kappa t} - 1}\right) \searrow 0$$

as  $|x| \rightarrow \infty$ ,

the Gaussian lower estimate does NOT hold.

It follows

$$\exp\left(-\frac{\kappa(y - xe^{\kappa t})^2}{e^{2\kappa t} - 1}\right)\Big|_{y=e^{\kappa t}x} = 1,$$

On the other hand, when  $y = e^{\kappa t}x$ ,

$|y - x| = |e^{\kappa t} - 1| \cdot |x| \rightarrow \infty$  as  $|x| \rightarrow \infty$ .

Hence, the Gaussian upper estimate does NOT hold.

## 7. A remark of Hölder continuity in time

Let  $T > 0$  and fix  $T$ .

Let  $p(s, x; t, y)$  be a non-negative Borel-measurable function s.t.

$$\int_{\mathbb{R}^d} p(s, x; t, \xi) d\xi = 1$$

$$\int_{\mathbb{R}^d} p(s, x; t, \xi) p(t, \xi; T, y) d\xi = p(s, x; T, y).$$

If  $p(s, x; t, y)$  is a transition probability density function of a Markov process, then the equations above hold.

## Proposition

Assume that

$$|p(t, x_1; T, y) - p(t, x_2; T, y)| \leq C(T - t)^{-d/2 - \beta} |x_1 - x_2|^\alpha$$
$$p(s, x; t, y) \leq \frac{C_+}{(t - s)^{d/2}} \exp\left(-\frac{\gamma_+ |x - y|^2}{t - s}\right).$$

Then, for  $s < t$

$$|p(t, x; T, y) - p(s, x; T, y)| \leq \tilde{C}(T - t)^{-d/2 - \beta} (t - s)^{\alpha/2}$$

where  $\tilde{C}$  is a constant depending on

$d, C, \alpha, C_+$  and  $\gamma_+$ .

## Proof.

$$\begin{aligned} & |p(t, x; T, y) - p(s, x; T, y)| \\ &= \left| p(t, x; T, y) \int_{\mathbb{R}^d} p(s, x; t, \xi) d\xi - \int_{\mathbb{R}^d} p(s, x; t, \xi) p(t, \xi; T, y) d\xi \right| \\ &\leq \int_{\mathbb{R}^d} |p(t, x; T, y) - p(t, \xi; T, y)| p(s, x; t, \xi) d\xi \\ &\leq \frac{C_+ C (T - t)^{-d/2 - \beta}}{(t - s)^{d/2}} \int_{\mathbb{R}^d} |\xi - x|^\alpha \exp\left(-\frac{\gamma_+ |\xi - x|^2}{t - s}\right) d\xi \\ &= \frac{C_+ C (T - t)^{-d/2 - \beta}}{(t - s)^{d/2}} \\ &\quad \times \int_{\mathbb{R}^d} \left(\frac{t - s}{\gamma_+}\right)^{\alpha/2} |\tilde{\xi}|^\alpha \exp(-|\tilde{\xi}|^2) \left(\frac{t - s}{\gamma_+}\right)^{d/2} d\tilde{\xi} \\ &= \tilde{C} (T - t)^{-d/2 - \beta} (t - s)^{\alpha/2}. \end{aligned}$$